

# Bayes estimators

Given  $\theta, X_1, \dots, X_n$  iid  $\sim f(x|\theta)$ .

prior  $g(\theta)$ .

joint distribution:  $f(x_1, \dots, x_n|\theta)g(\theta)$

posterior  $g(\theta|x_1, \dots, x_n)$ .

Estimator  $T = h(X_1, \dots, X_n)$ .

Loss:  $L(t, \theta)$

Risk:  $R(T, \theta) = E_T L(T, \theta)$

Bayes risk of  $T$  is

$$R(T, g) = \int_{\Theta} R(T, \theta)g(\theta)d\theta.$$

$T$  is Bayes estimator if

$$R(T, g) = \inf_T R(T, g).$$

To minimize Bayes risk, we only need to minimize the conditional expected loss given each  $x$  observed.

# Binomial model

$$X \sim \text{bin}(x|n, \theta)$$

prior  $g(\theta) \sim \text{unif}(0, 1)$ .

posterior  $g(\theta|x) \sim \text{beta}(x + 1, n - x + 1)$ .

Beta distribution:

$$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}, 0 \leq x \leq 1, E(X) = \frac{\alpha}{\alpha+\beta}.$$

For squared error loss, the Bayes estimator is

posterior mean  $\frac{x+1}{n+2}$ .

If  $X$  is a random variable, choose  $c$

$$\min E(X - c)^2.$$

$$c = E(X)$$

# Prior distributions

Where do prior distributions come from?

- \* a prior knowledge about  $\theta$
- \* population interpretation—(a population of possible  $\theta$  values).
- \* mathematical convenience (conjugate prior)

conjugate distribution— the prior and the posterior distribution are in the same parametric family.

# Conjugate prior distribution

## **Advantages:**

- \* mathematically convenient
- \* easy to interpret
- \* can provide good approximation to many prior opinions (especially if we allow mixtures of distributions from the conjugate family)

## **Disadvantages:**

may not be realistic

# Binomial model

$$X|\theta \sim \text{bin}(n, \theta)$$

$$\theta \sim \text{beta}(\alpha, \beta).$$

Beta distribution:

$$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1},$$

$$0 \leq x \leq 1, E(X) = \frac{\alpha}{\alpha+\beta}.$$

$$g(\theta|x) = \text{beta}(x + \alpha, n - x + \beta).$$

$$\text{posterior mean is } \hat{\theta} = \frac{x+\alpha}{n+\alpha+\beta}.$$

## Binomial model: continued

Under loss function  $L(t, \theta) = \frac{(t-\theta)^2}{\theta(1-\theta)}$ .

We need to minimize

$$\begin{aligned} g(t) &= E_{\theta|X} L(t, \theta) = E_{\theta|X} \frac{(t-\theta)^2}{\theta(1-\theta)} \\ &= t^2 E_{\theta|X} \frac{1}{\theta(1-\theta)} - 2t E_{\theta|X} \frac{1}{1-\theta} + E_{\theta|X} \frac{\theta}{1-\theta}. \end{aligned}$$

It is a quadratic function of  $t$ .

The minimizer is

$$t^* = \frac{E_{\theta|X} \frac{1}{1-\theta}}{E_{\theta|X} \frac{1}{\theta(1-\theta)}} = \frac{\alpha + x - 1}{\alpha + \beta + n - 2}.$$

## exercise

Let  $X_1, \dots, X_n$  be iid Poisson ( $\lambda$ ), and let  $\lambda$  have a gamma ( $\alpha, \beta$ ) distribution.

- 1). Find the posterior distribution of  $\lambda$ .
- 2). Calculate the posterior mean and variance.
- 3). Find the Bayes estimator under the loss  $L(t, \lambda) = \frac{(t-\lambda)^2}{\lambda}$ .



# Solutions

1). The joint distribution of  $X_1, \dots, X_n$  given  $\lambda$  is

$$p(x_1, \dots, x_n | \lambda) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} = \frac{e^{-n\lambda} \lambda^{\sum_i x_i}}{\prod_i x_i!}$$

the posterior distribution of  $\lambda$  is

$$g(\lambda | x_1, \dots, x_n) \propto p(x_1, \dots, x_n | \lambda) g(\lambda)$$

$$\propto \frac{e^{-n\lambda} \lambda^{\sum_i x_i}}{\prod_i x_i!} \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$

$$\propto \lambda^{\sum_i x_i + \alpha - 1} e^{-(\beta+n)\lambda} \propto \text{gamma}(\sum_i x_i + \alpha, \beta + n).$$

2). The posterior mean of  $\lambda$  is  $E(\lambda | x_1, \dots, x_n) = \frac{\sum_i x_i + \alpha}{\beta + n},$

and the posterior variance is  $Var(\lambda | x_1, \dots, x_n) = \frac{\sum_i x_i + \alpha}{(\beta + n)^2}.$

3). For simplicity, write  $x = (x_1, \dots, x_n)$ , we need to minimize

$$E_{\lambda|x} L(t, \lambda) = E_{\lambda|x} \frac{(t-\lambda)^2}{\lambda} = E_{\lambda|x} \left( \frac{1}{\lambda} t^2 - 2t + \lambda \right)$$

which is a quadratic function of  $t$ . The minimizer is

$$t^* = \frac{1}{E_{\lambda|x} \frac{1}{\lambda}}.$$

$$\text{Note } E_{\lambda|x} \frac{1}{\lambda} = \int \frac{1}{\lambda} \frac{(\beta+n)^{\sum_i x_i + \alpha}}{\Gamma(\sum_i x_i + \alpha)} \lambda^{\sum_i x_i + \alpha - 1} e^{-(\beta+n)\lambda} d\lambda =$$

$$\frac{\Gamma(\sum_i x_i + \alpha - 1) (\beta+n)^{\sum_i x_i + \alpha}}{\Gamma(\sum_i x_i + \alpha) (\beta+n)^{\sum_i x_i + \alpha - 1}} = \frac{\beta+n}{\sum_i x_i + \alpha - 1}.$$

$$\text{and the Bayes estimator is } t^* = \frac{\sum_i x_i + \alpha - 1}{\beta + n}.$$

Let  $\Theta$  consists of two points,  $\Theta = \{\frac{1}{3}, \frac{2}{3}\}$ .

Let  $\mathcal{A}$  be the real line and  $L(\theta, a) = (\theta - a)^2$ .

A coin is tossed once and the probability of head is  $\theta$ . Consider the set of nonrandomized decision rules which are functions from the set  $\{H, T\}$  to  $\mathcal{A}$ :  $d(H) = x, d(T) = y$ .

Find the Bayes rule with respect to the prior distribution giving prob.  $1/2$  to  $\theta = 1/3$  and prob.  $1/2$  to  $\theta = 2/3$ .

## Exercise

Suppose the conditional distribution of given  $\theta$  is  $X|\theta \sim \text{bin}(3, \theta)$ , and the prior distribution of  $\theta$  is uniform on  $(0, 1)$ . i.e.,  $g(\theta) = 1, 0 \leq \theta \leq 1$ .

Suppose  $X = 2$  is observed.

- 1). Find the joint distribution of  $X$  and  $\theta$ .
- 2). Find the probability  $P(X = 2)$ .
- 3). Derive the conditional distribution of  $\theta$  given  $X = 2$ .
- 4). Find the conditional mean and variance of  $\theta$  given  $X = 2$ .